

Identifying and Estimating Dynamical Covariance Matrices with Hierarchical Rank Structure

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Covariance Matrices in Dynamical Systems Modeling

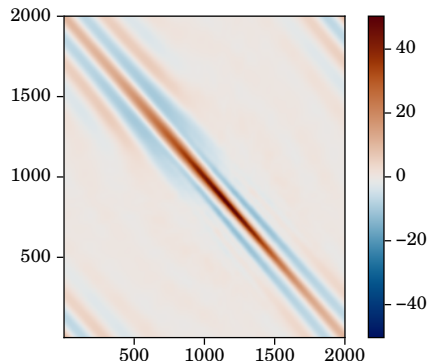
- Consider a dynamical system in n dimensions:

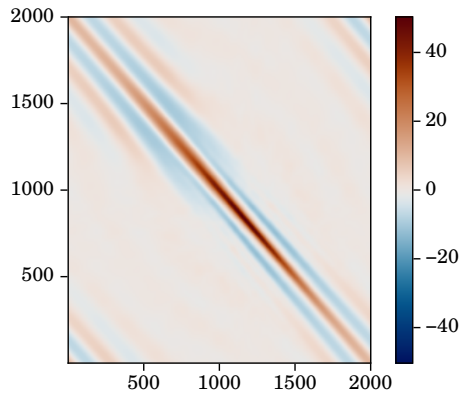
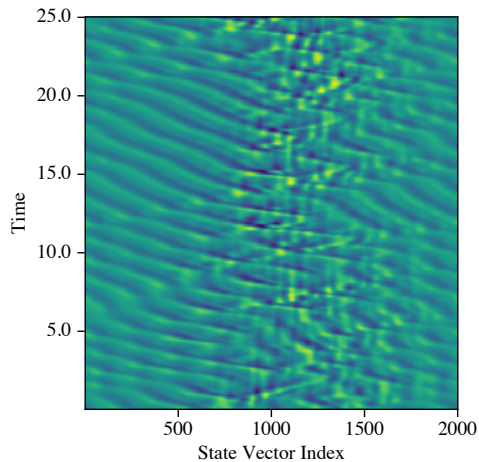
$$\mathbf{x} \in \mathbb{R}^n, \quad \frac{d\mathbf{x}}{dt} = f(\mathbf{x}, t).$$

- **Goal:** to estimate the covariance matrix of a measure associated with the dynamics. For example:
 - Model reduction: covariance of the *invariant measure* reveal the appropriate subspace for POD.
 - Data assimilation: covariance of a *prior measure* represents uncertainty about the state, and spreads information from observed variables to unobserved ones.
- We begin with an *ensemble* or *sample* covariance estimate:

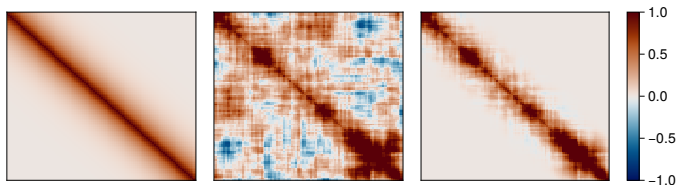
$$\hat{\Sigma} = \frac{1}{s-1} \sum_{i=1}^s (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})^T,$$

where $\mathbf{x}_1, \dots, \mathbf{x}_s$ are an ensemble of model states.





Sampling Errors and Localization



- Sample size may be *much* smaller than phase space dimension; $s \ll n$.
- In this case, $\hat{\Sigma}$ will be unacceptably noisy.
- **Localization** [4] regularizes $\hat{\Sigma}$ by attenuating “unphysical” correlations:

$$\Sigma = \mathbf{L} \circ \hat{\Sigma},$$

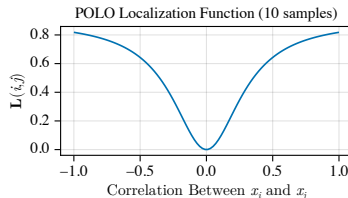
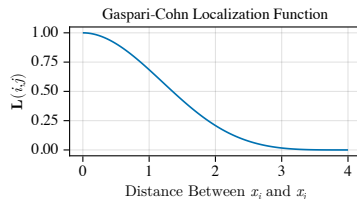
where \circ = element-wise product, \mathbf{L} = a symmetric localizing matrix in $[0, 1]^{n \times n}$.

Localization Functions

- **Distance-based localization:** correlations are attenuated based on spatial separation between variables.
 - Admits a data-sparse matrix representation.
 - Encodes *a priori* physical structure.
 - Careful tuning needed for localization radius.
- **Prior optimal localization (POLO):** [6] small empirical correlations are attenuated more than large ones.

$$\mathcal{L}(\rho) = \frac{(s-1)\rho^2}{1+s\rho^2}.$$

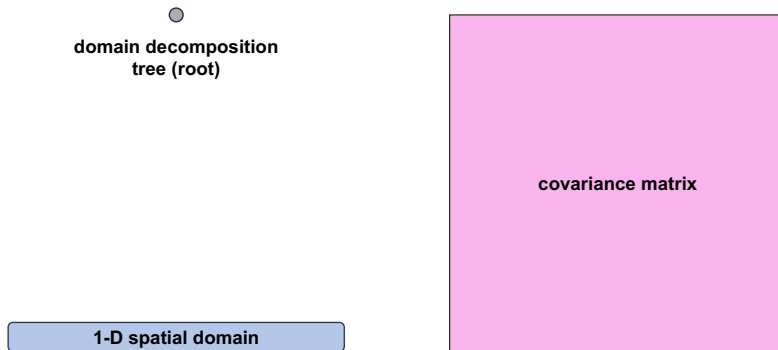
- No tuning needed.
- Adapts to sample size.
- **Optimal for multivariate Gaussian samples [6].**
- **Expensive to construct for high-dimensional systems.**



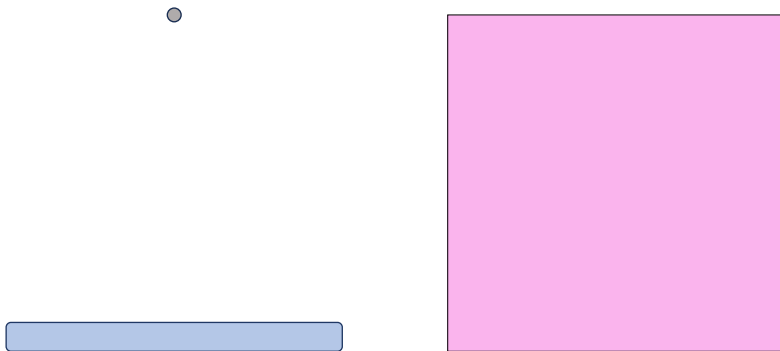
Our Contributions

- **The problem we address:** POLO localization does not produce a data-sparse matrix.
 - ...at least, not without $\mathcal{O}(n^2)$ work to create sparsity!
- **Our solution:** represent the POLO estimator as a hierarchically rank-structured matrix [3].
 - Spatial domain is recursively partitioned into nested subdomains.
 - At each level, admissible blocks (correlations between sufficiently far-apart subdomains) are compressed to low-rank form.
- **The challenge we face:** we must create a hierarchically rank structured representation by only looking at our samples.
 - We are not allowed to form the dense-matrix POLO estimator first!

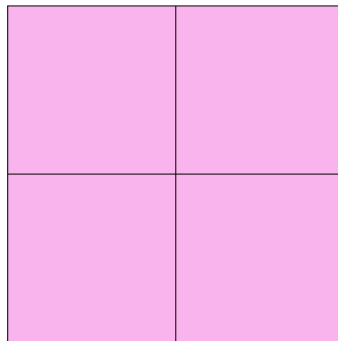
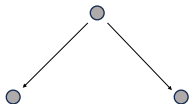
Compression Using Hierarchical Rank Structure



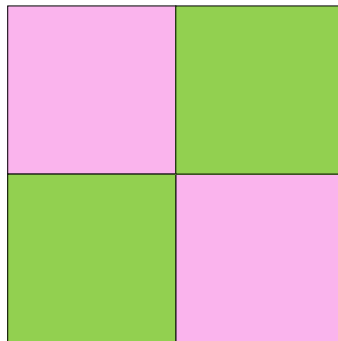
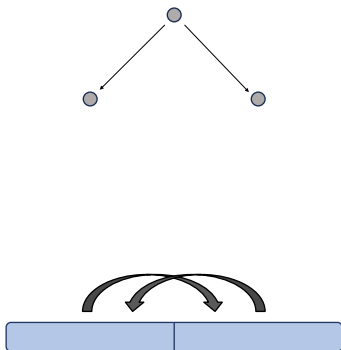
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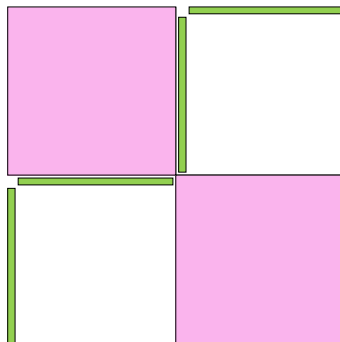
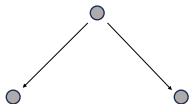
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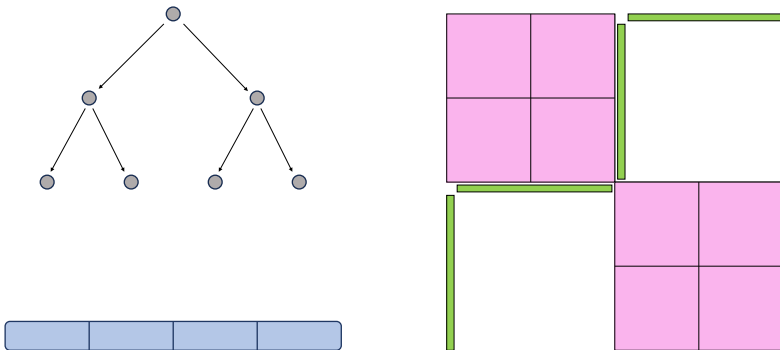
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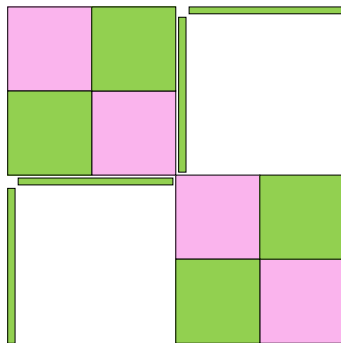
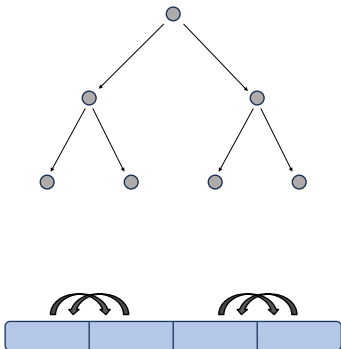
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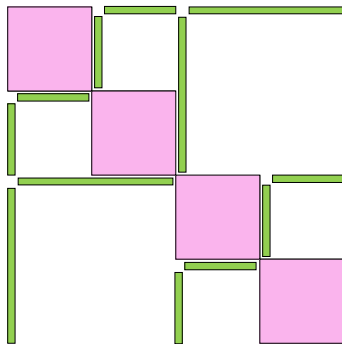
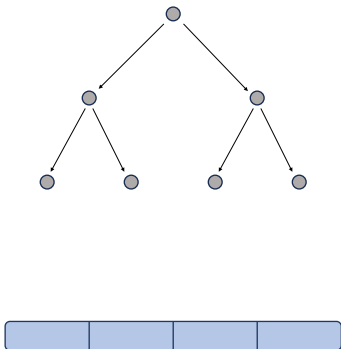
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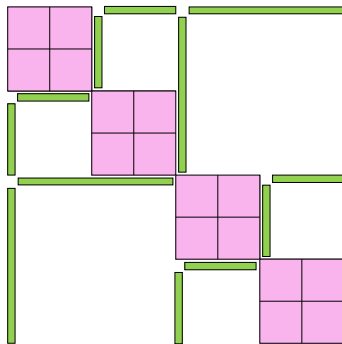
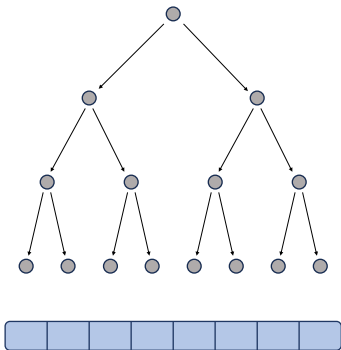
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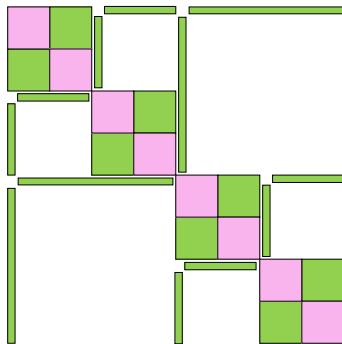
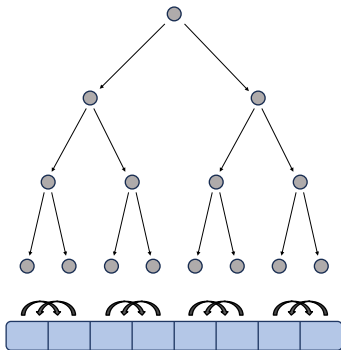
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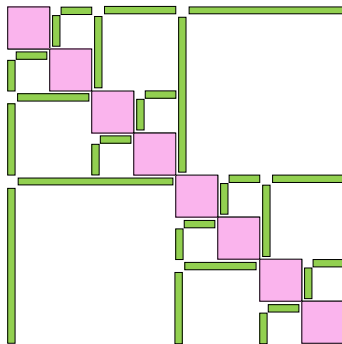
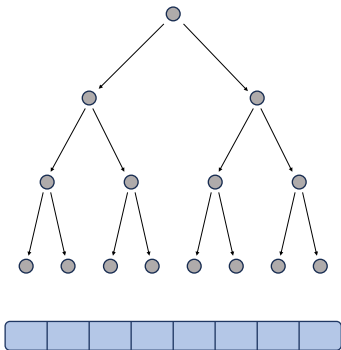
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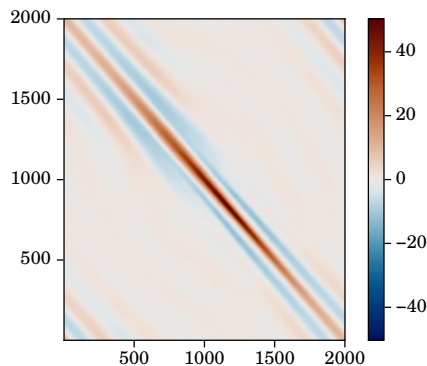


Compression Using Hierarchical Rank Structure

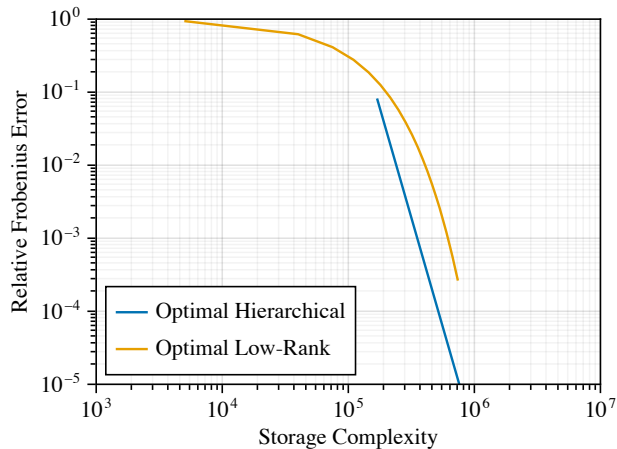
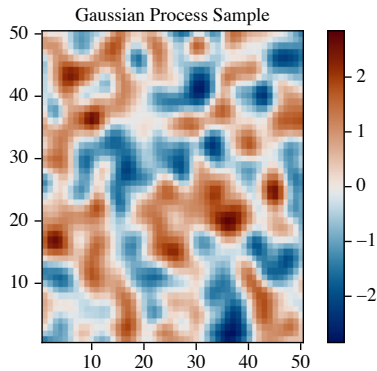


Is Hierarchical Rank Structure Appropriate Here?

- **Informally:** hierarchical rank structure is appropriate *when long-distance correlations vary more smoothly than short-distance correlations*.
- **More Formally:** kernels that are *asymptotically smooth* [3] (correlations and their derivatives decay algebraically away from the diagonal) have hierarchical rank structure.
- Many details to consider, such as:
 - What is the procedure for recursively partitioning space?
 - What is the criterion for determining which blocks are admissible (can be compressed)?
- **Examples of hierarchical rank structure:**
 - Greens functions associated with elliptic operators.
 - **Diffusion-based covariance models used in numerical weather prediction** [8].



Hierarchical Rank Structure vs Low Rank Structure



Estimating Admissible Blocks: Overview

- **Notation:** $\hat{\Sigma}$ = the unlocalized covariance matrix, $(\mathcal{X}, \mathcal{Y})$ = an admissible subdomain pair, and

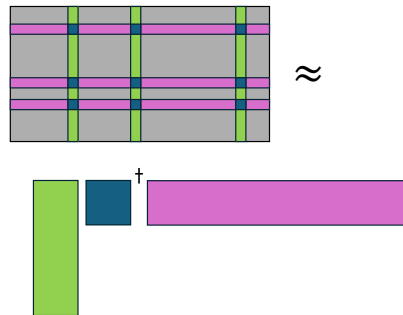
$$\mathcal{L}(\rho) = \frac{(s-1)\rho^2}{1+s\rho^2} = \text{POLO localizer for } s \text{ samples.}$$

- **Goal:** form an approximate low-rank factorization of a localized admissible block,

$$\Sigma(\mathcal{X}, \mathcal{Y}) = \mathcal{L}(\hat{\Sigma}(\mathcal{X}, \mathcal{Y})) \approx \mathbf{L}_{\mathcal{X}} \mathbf{L}_{\mathcal{Y}}^T,$$

using as few entry evaluations in $\Sigma(\mathcal{X}, \mathcal{Y})$ as possible.

- **Approach:** a generalized Nyström approximation [7]. The challenge is to select good skeleton rows/columns without looking “too hard” at $\Sigma(\mathcal{X}, \mathcal{Y})$.



Column/Row Selection via Proxy Points

- We want to choose a skeleton column set $\mathcal{S} \subseteq \mathcal{Y}$, $|\mathcal{S}| = r \ll |\mathcal{Y}|$, such that

$$\sigma_{\min}(\mathbf{\Sigma}(\mathcal{X}, \mathcal{S})) \approx \max_{\mathcal{T} \subseteq \mathcal{Y}, |\mathcal{T}|=r} \sigma_{\min}(\mathbf{\Sigma}(\mathcal{X}, \mathcal{T})).$$

- **Column-pivoted QR (CPQR)** factorization provides a good solution,

$$[\mathbf{\Sigma}(\mathcal{X}, \mathcal{S}) \quad \mathbf{\Sigma}(\mathcal{X}, \mathcal{Y} \setminus \mathcal{S})] = \mathbf{Q} [\mathbf{R}_1 \quad \mathbf{R}_2],$$

but requires evaluating all the entries in $\mathbf{\Sigma}(\mathcal{X}, \mathcal{Y})$.

- **Proxy point method:** select a set of **proxy points** [9] $\mathcal{P} \subseteq \mathcal{X}$ with $|\mathcal{P}| \ll |\mathcal{X}|$. Pivot on $\mathbf{\Sigma}(\mathcal{P}, :)$ instead:

$$[\mathbf{\Sigma}(\mathcal{P}, \mathcal{S}) \quad \mathbf{\Sigma}(\mathcal{P}, \mathcal{Y} \setminus \mathcal{S})] = \mathbf{Q} [\mathbf{R}_1 \quad \mathbf{R}_2].$$

- How to choose proxy points without looking at $\mathbf{\Sigma}(\mathcal{X}, \mathcal{Y})$?

Proxy Point Selection

- Proxy points should correspond to rows that capture the rank structure of $\text{Row } \Sigma(\mathcal{X}, \mathcal{Y})$.
- Use **CPQR** on rows of $\hat{\Sigma}(\mathcal{X}, \mathcal{Y})$ (unlocalized).
- If $\mathbf{Z}_{\mathcal{X}}$, $\mathbf{Z}_{\mathcal{Y}}$ are centered, normalized samples restricted to \mathcal{X} and \mathcal{Y} , then

$$\hat{\Sigma}(\mathcal{X}, \mathcal{Y}) = \mathbf{Z}_{\mathcal{X}} \mathbf{Z}_{\mathcal{Y}}^T = (n_{\mathcal{X}} \times s) \cdot (s \times n_{\mathcal{Y}}),$$

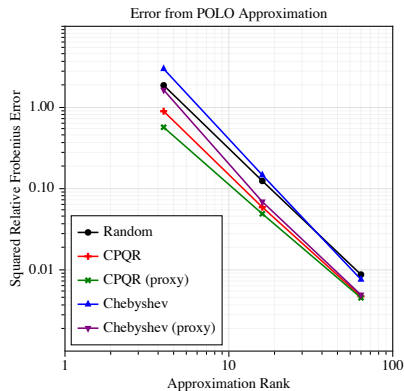
where $s \ll \min\{n_{\mathcal{X}}, n_{\mathcal{Y}}\}$.

Lemma

CPQR factorization on $\mathbf{A}_1 = \mathbf{Z}_{\mathcal{Y}} \mathbf{Z}_{\mathcal{X}}^T(:, l)$ yields an identical column permutation as CPQR factorization on

$$\mathbf{A}_2 = (\mathbf{Z}_{\mathcal{Y}}^T \mathbf{Z}_{\mathcal{Y}})^{1/2} \mathbf{Z}_{\mathcal{X}}^T(:, l).$$

Note that \mathbf{A}_1 is $n_{\mathcal{Y}} \times n_{\mathcal{X}}$, while \mathbf{A}_2 is only $s \times n_{\mathcal{X}}$.



Full Estimation Procedure

Algorithm 1 Estimating Low-Rank Factors Of An Admissible Block

- 1: **input:** $\mathbf{Z}_{\mathcal{X}} \in \mathbb{R}^{n_{\mathcal{X}} \times s}$, $\mathbf{Z}_{\mathcal{Y}} \in \mathbb{R}^{n_{\mathcal{Y}} \times s}$ (centered and normalized sample vectors on \mathcal{X} and \mathcal{Y})
- 2: **input:** $k, p \geq 1$ (admissible block rank and skeletonization rank)
- 3: $\mathcal{P}_1 \leftarrow \text{CPQRColumnSelect}((\mathbf{Z}_{\mathcal{Y}}^T \mathbf{Z}_{\mathcal{Y}})^{1/2} \mathbf{Z}_{\mathcal{X}}^T, p)$ *# selecting p skeleton columns*
- 4: $\mathcal{S}_{\mathcal{Y}} \leftarrow \text{CPQRColumnSelect}(\boldsymbol{\Sigma}(\mathcal{P}_1, :), p)$
- 5: $\mathcal{P}_2 \leftarrow \text{CPQRColumnSelect}((\mathbf{Z}_{\mathcal{X}}^T \mathbf{Z}_{\mathcal{X}})^{1/2} \mathbf{Z}_{\mathcal{Y}}^T, p)$ *# selecting p skeleton rows*
- 6: $\mathcal{S}_{\mathcal{X}} \leftarrow \text{CPQRColumnSelect}(\boldsymbol{\Sigma}(:, \mathcal{P}_2)^T, p)$
- 7: $\mathbf{L}_1 \mathbf{L}_2^T \leftarrow \boldsymbol{\Sigma}(:, \mathcal{S}_{\mathcal{Y}}) \boldsymbol{\Sigma}(\mathcal{S}_{\mathcal{X}}, \mathcal{S}_{\mathcal{Y}})^{\dagger} \boldsymbol{\Sigma}(\mathcal{S}_{\mathcal{X}}, :)$ *# rank- p generalized Nyström approximation*
- 8: $\mathbf{Q}_1, \mathbf{R}_1 \leftarrow \text{ThinQR}(\mathbf{L}_1)$, $\mathbf{Q}_2, \mathbf{R}_2 \leftarrow \text{ThinQR}(\mathbf{L}_2)$ *# optimal reduction to rank k*
- 9: $\mathbf{U}, \boldsymbol{\Gamma}, \mathbf{V} \leftarrow \text{SVD}(\mathbf{R}_1 \mathbf{R}_1^T)$
- 10: $\mathbf{B}_1 \leftarrow \mathbf{Q}_1 \mathbf{U}(:, 1:r) \boldsymbol{\Gamma}(1:r, 1:r)^{1/2}$
- 11: $\mathbf{B}_2 \leftarrow \mathbf{Q}_2 \mathbf{V}(:, 1:r) \boldsymbol{\Gamma}(1:r, 1:r)^{1/2}$
- 12: **return** $\mathbf{B}_1, \mathbf{B}_2$

Test Case 1: “Storm Track” Dynamics

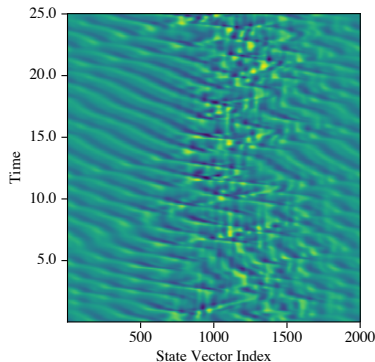
- A modification of “model II” from Lorenz [5]. Like the Lorenz ‘96 model, but:
 - admits waves much larger than the grid spacing, and
 - has a “stable” region of strong damping and a “chaotic” region of weak damping.

Based off a system in [1].

- **Domain:** 2000 grid points in 1D with periodic boundary conditions.
- **Partition:** recursive bisection until domain has at most 10 grid points.
- **Admissibility criterion:**

$$\min\{\ell(\mathcal{X}), \ell(\mathcal{Y})\} \leq d(\mathcal{X}, \mathcal{Y}),$$

where $\ell(\cdot)$ = domain length, and
 $d(\mathcal{X}, \mathcal{Y}) = \inf_{x \in \mathcal{X}, y \in \mathcal{Y}} |x - y|$.



Test Case 2: 2D Gaussian Process

■ “Ground-truth” covariance:

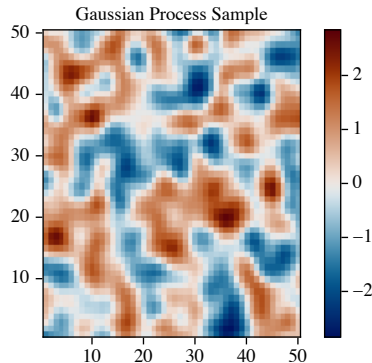
$$\Sigma_0((x_i, y_i), (x_j, y_j)) = \exp\left(-\frac{(x_i - x_j)^2 + (y_i - y_j)^2}{2\sigma^2}\right),$$

where $\sigma = 0.05$ and (x_i, y_i) are nodes of a 50×50 uniform grid on $[0, 1] \times [0, 1]$.

- Samples in \mathbb{R}^{2500} drawn i.i.d. from $\mathcal{N}(\mathbf{0}, \Sigma)$.
- **Partition:** recursive bisection of rectangles along longer axis, until max. sidelength is below 0.2.
- **Admissibility criterion:**

$$\min\{\ell(\mathcal{X}), \ell(\mathcal{Y})\} \leq d(\mathcal{X}, \mathcal{Y}),$$

where $\ell(\cdot) = \max$ sidelength of rectangle, and $d(\mathcal{X}, \mathcal{Y}) = \inf_{\mathbf{x} \in \mathcal{X}, \mathbf{y} \in \mathcal{Y}} \|\mathbf{x} - \mathbf{y}\|_2$.



Test Case 3: 2D Quasigeostrophic Turbulence

- **Quasigeostrophic flow** approximates the motion of a rotating fluid where Coriolis and pressure-gradient forces are nearly in balance [2].
- Commonly used as a simplified model of atmospheric flow.
- Covariance has *mild spatial inhomogeneity*.
- **Domain:** 128×128 grid on a 2D square with periodic boundary conditions.
- **Partition:** bisecting rectangles until longest side spans at most 10 gridpoints.
- **Admissibility Criterion:** same as before.
- Simulated with code from <https://github.com/jswhit/sqgturb>.

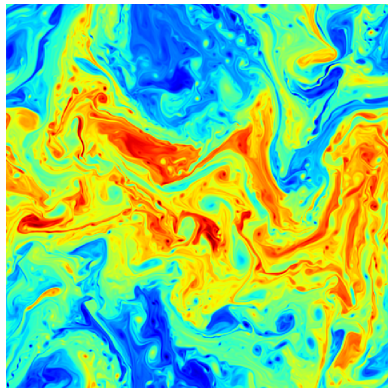
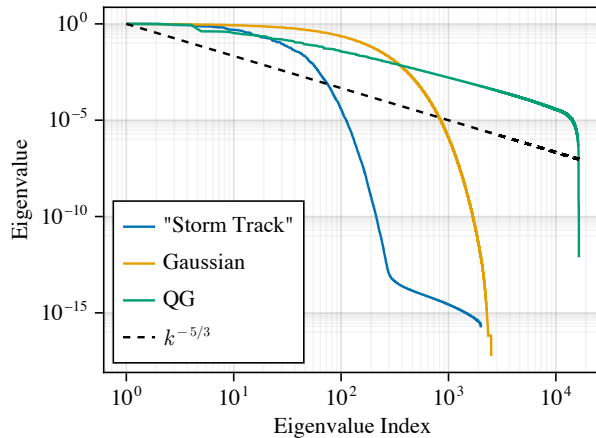
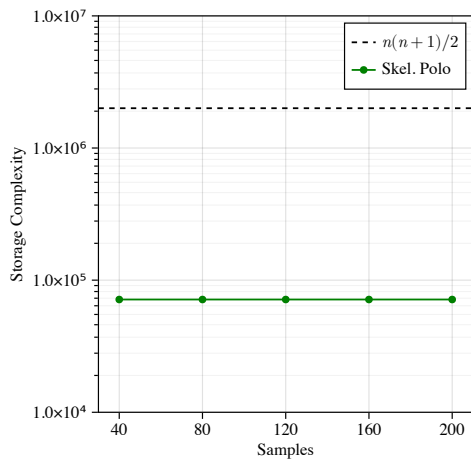
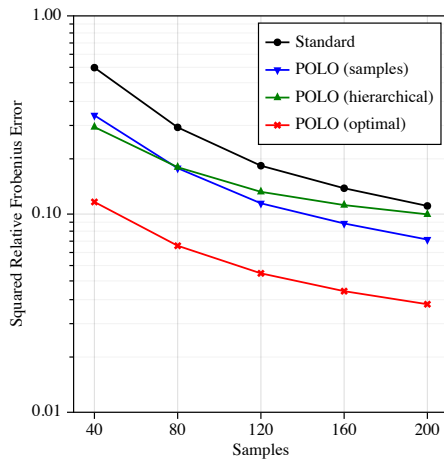


Figure: from
<https://github.com/jswhit/sqgturb>.

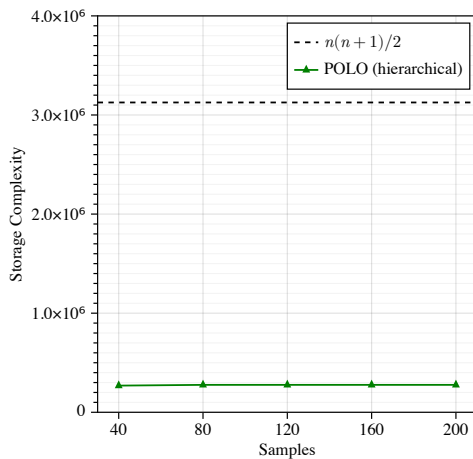
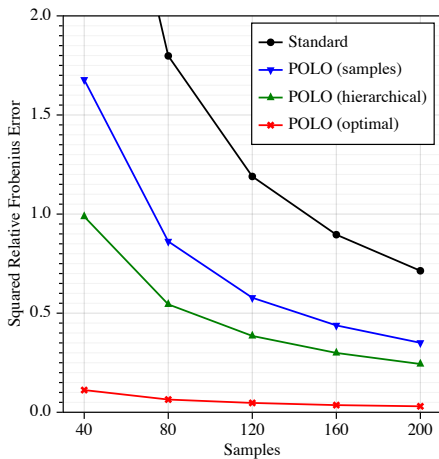
Problem Difficulty



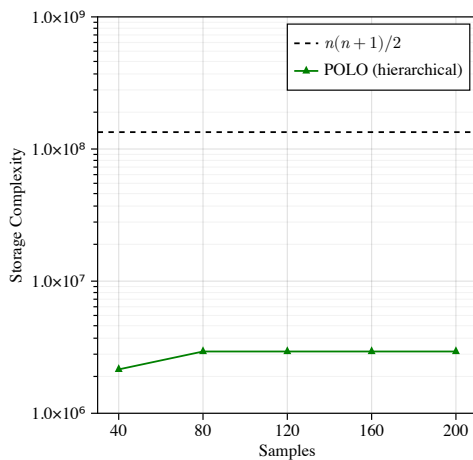
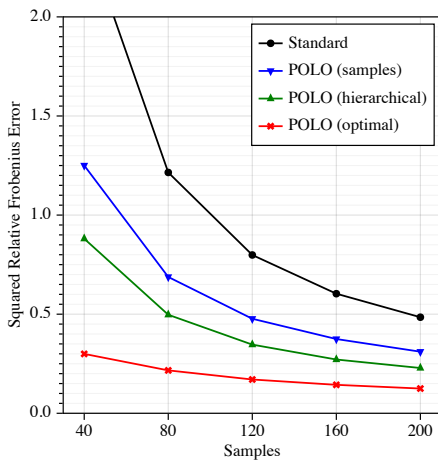
Results: “Storm Track” Dynamics



Results: Gaussian Process







Results: Quasigeostrophic Turbulence







Conclusions

- Localization is an essential aspect of covariance matrix estimation when very few samples are available.
- The POLO localizer has optimality properties and requires no tuning, but expensive to construct and store.
- **We have developed:** *a data-sparse, efficiently constructable covariance estimator that corresponds to a hierarchically rank-structured approximation of POLO localization.*
- **Next steps:**
 - 1 More complex test cases, including 3D spatial domains.
 - 2 **Positive-definite estimators**, which POLO itself is not.
 - 3 Testing performance in model reduction and data assimilation problems.

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